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Do Peer Firms Affect Trade Credit Policies?

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Abstract— The paper examines the effect of peer firm decisions with regard to financial policies. Using data on NSE listed firms of oil and Petro chemical industry for the period 2013 to 2022, we document strong evidence of peer influence on trade credit provision. The result results are obtained by 2SLS after using alternative trade measures and addressing the concerns of endogeneity. The basic vital ratios are also addressed to get an overall understating of the industry.

Index Terms— Trade Credit; Accounts Receivable, Firm Behaviour, Peer Effects, Mimicking

I. INTRODUCTION

Trade credit is one among the most important source of short-term finance for firms in developed countries ([18]; [1]) Across the Globe it is held responsible for global trade and commerce in excess of US\$25 trillion [11]. Younger and developing firms find trade credit as a natural source of capital to both financial and non- financial firms [2]. Firms trade action depending on their peers has not been much explored in Indian context unlike other developing nation. There has been a many key studies which underline the importance of peer influence in deciding the trade credit policies of corporates. ([6]; [1];[9]).

Trade receivables have been put on a common practise and the relevance of this policy for the corporate sector, prior to theoretical research has been highlighted by various financing and no financing firms for trade credit provision. There have been many research documenting firm-specific factors determining the use of accounts receivable among corporates ([15];[16];[18]). There is evidence that trade credit provision is affected by many other factors like macro-financial shocks [3] financial crises [7], and national culture [5]. Most research implicitly assumes that firms manage trade credit policies in isolation and independently without considering the actions and characteristics of their rival firms in the same industry and supply similar products. As such, the question of whether a firm's trade credit decisions are affected by those of its peers is still unanswered ([12];[6]; [9][18]).

II. DATA & METHODOLOGY

The data used in this study were obtained from the Prowess database. The initial sample includes all listed Oil & Gas companies in the National Stock Exchange covered by prowess database for the period from 2013 to 2022. Unlike many developed economics such as France, Germany and Japan where the banking system dominates credit allocation, in India the equity market dominates the same.[4]. According to [9], the ratio of trade receivables to total assets of UK firms is 20.47 per cent. Financial firms such as banks and insurance were excluded because they have different accounting requirements [4]. Moreover, firm-years with anomalies in their accounts such as negative values in assets, sales, trade receivable, trade payable and fixed assets were removed. Also, firms missing more than five years' amount of information and duplicate values were excluded.[8]). Oil & Gas is a prominent sector in the manufacturing sector consisting of 20 companies in India. The average market capitalization of the industry is 122927.525 Crores., with a Net worth amounting to 52364.63 Crores. The major giants of the industry are Reliance Industries & ONGC with a market capitalization of 17,45,813 and 2,13,236 crores [5] respectively. Four companies have been excluded due to insufficiency of data and for being younger than 10 years in 2022. As a result, 16 firms are left with 1800 firm-year observations. (Table 1.)

Table 1. List of	Companies
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	Sector	Number of companies listed	No. of Companies selected	List of Company exempted
1.	Oil & Gas	21	16	 Asian Energy services ltd Mahanagar Gas Ltd Deep Energy Resources Ltd Gujarat Gas Ltd

Following recent research on peer effects ([14];[6]; [9])., we use the following model to examine the impact of peer firm behavior on a firm's trade credit policy where the subscripts *i*, *j*, and *t* denote firm, industry, and year, respectively; $y_{i,j,t}$ is the dependent variable capturing firm *i*'s trade credit provision; y^-_{-ijt} is the average trade receivables of Oil& natural Gas industry based peers, excluding firm *i*; $X_{ij,t-}$ *i* are one-year lagged control variables; $X^-_{-ij,t-1}$ are one-year lagged averages of the same control variables for peer firms, excluding firm *i* (i.e., the peer averages); μ_t is year fixed effects; and ϵ_{ijt} is the error term.

We include the year fixed effects to control for potential changes in trade credit provision over time, including periods of crises ([12;[13]). In our regression analysis, we are mainly interested in the two coefficients β and γ' , which capture peer



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effects through peer firms' actions and characteristics, respectively. Using OLS to estimate Model (1) may lead to a reflection problem [15] When one regresses the outcome variable on its group averages a specific type of endogeneity may arise [15];[16]). In our analysis, the regressing an individual firm's trade credit provision on its peer averages may not be justifiable one due to the endogenous selection of firms into peer groups ([14]). Omitted variable bias arise due to omitted factors that influence both the firm's and its peers' trade credit decisions. As reflection problem to poses a major challenge research studying peer effects in corporate finance ([15]; [14]; [10]. For identification purposes, we substitute peer idiosyncratic stock returns, this variable is satisfies both the relevance and exclusion conditions of a valid instrument for peer effects [15]. To calculate the instrument, peer idiosyncratic stock returns, we first estimate the expected stock return using the [7] four-factor model, as follows:

$$r_{ijt} = \alpha_i + \beta_{1i}(rm_t - rf_t) + \beta_{2i}SMB_t + \beta_{3i}HML_t + \beta_{4i}MOM_t + 5_{ijt} (2)$$

where r_{ijt} captures the stock return of firm *i*, in industry *j*, and over month *t*. The four actors in the model are the excess market return $(rm_t - rf_t)$, the small minus big portfolio (SMB_t) returns (size factor), the high minus low (HML_t) portfolio returns (value factor) and momentum (MOM). We obtain stock returns data from Fama and French factors (i.e., market returns, *SMB*, MOM and *HML*) from Kenneth French's database. We estimate expected stock returns by using a monthly rolling regression approach to estimate Equation (2). Using the estimated coefficients from Equation (2), we compute the expected stock returns as follows:

$r_{ijt} = \alpha^{\gamma} + ^{1}i(rm_{t} rf_{t}) ^{2}iSMB_{t} + \beta^{3}iHML_{t} ^{2}iSMB_{t} + \beta^{4}iMOM$ (3)

We then convert the monthly returns data into annual returns to match with our firm- year data. In the next step, we calculate idiosyncratic stock returns as the difference between the observed and expected stock returns, as follows:

 $Idio_{-ijt} = r_{ijt} - r_{ijt}$

(4)

We compute our instrument, $Idio_{-ij,t-1}$, by taking one-year lagged average peer idiosyncratic stock returns, where j represents the peers of firm i.

Finally, we use $Idio_{-ij,t-1}$ as an instrument for peer influence in our 2SLS regressions. Our approach involves estimating Equations (5) and (6) in two stages,

$$y-ijt = \alpha + \beta I I dio-ij, t-1 + \beta 2 \overline{X}^{-} - ij, t-1 + \beta 3 X ij, t-1 + \beta 4 \mu t + \xi ijt$$
(5)
$$y_i, j, t = \alpha + \beta y^{-} - ijt + \gamma' \overline{X}^{-} - ij, t-1 + \lambda' X ij, t-1 + \frac{\partial^2}{\mu} \mu t + \epsilon ijt$$
(6)

In the first stage, we estimate the reduced-form model (Equation (5)), where we regress the endogenous variable, peer average trade receivables, on the instrument, $Idio_{-ijt-1}$, as well as the controls, namely, the firm-specific characteristics, average peer characteristics, and year fixed effects. We then obtain the fitted values for peer average trade credit, $_y^{-}_{i}$, *j*, *t*, use them in the second stage (Equation (6)) to estimate the peer effects, β , in trade credit policy.

III. DATA & INTERPRETATION

Table 2 provides the summary statistics of oil & Gas industry's firm specific characteristics. Accounts receivable represents a significant portion amounting to 19% of the current asset in the oil & Gas industry, The maximum representation is around 70% with a standard deviation of 14%. While, the mean accounts receivable represents only 1% (0.64) of total assets. The mean net operating cycle is 92 days, standard deviation of 215 days depicting high variation across firm in the industry. Debtors' turnover means to 24 times with a median of 20. is noteworthy that means are all bigger than medians. This indicates that mean values are influenced by big outliers. Table.3 To check whether there are significant differences between the debtors turnover within the sector. An ANOVA test of equality of means is undertaken. This test allows us to verify the null hypothesis, which indicates that the means of debtor's turnover of oil and gas companies are equal, against the alternative, which indicates that there are differences between the means of companies within the sector. The upper part of Table 3 shows the results of the hypothesis, we obtained a value of the contrast statistic of 87 .06 and a p-value of 0.00. With these data, at the 0.05 significance level we reject the null hypothesis that the Omeans for the different companies in the sector are equal. Therefore, we accept the alternate hypothesis that there is significant differences in the of debtor's turnover in the within the Oil and Gas sector. Across the period of 10 years the average debtor days is 27 days, analysing a trend for the first three years debtor days in table 4 shows low variation. The last three years also recorded relatively stable debtor days ranging from 23 to 32 days. Sales for the year 2020 recorded a negative growth but the quantum of variation is not that substantial. Aban offshore Ltd have been excluded from the analysis as its rate of turnover is much higher than the industry leaders. Besides having the longest debtor days in the whole sector, the company recorded high fluctuation in the Debtors days of around averaging to 275 days in the last three years is represented on Table: 4.

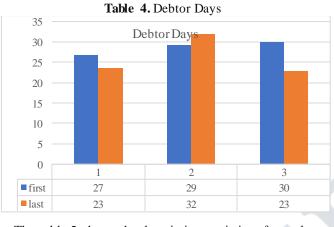
Table 2. Descriptive statistics

	Table 2. Descriptive statistics							
		REC_TO		DEBTORS				
	REC_TO_CA	ASSETS	NOPC	TURNOVER				
Mean	18.77	0.06	91.54	24.44				
Median	14.33	0.05	46.31	20.76				
Maximum	69.92	0.44	920.50	82.54				
Minimum	0.94	0.01	-518.07	0.24				
Std. Dev.	13.66	0.06	215.97	18.56				
Observations	150	150	150	150				



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Table 3: Test for Equality of Means of DEBTORS_TURNOVER					
Method	df	Value	Probability		
Anova F-test	(14, 135)	87.06006	0.0000		
Welch F-test*	(14, 50.1116)154.3371	0.0000		
*Test allows for unequal c	cell variances				
Analysis of Variance					
Source of Variation	df	Sum of Sq.	M ean Sq.		
Between	14	46208.44	3300.603		
Within	135	5118.092	37.91179		
Total	149	51326.53	344.4734		



The table 5 shows the descriptive statistics of petroleum industry. The average stock return of petroleum industry is (-) 4% (-0.0365). Table. 6 represent Augmented Dicky-Fuller test determines the series is stationary (p value 0.00). Table No.7 depicts determines the panel characteristics whether we have a pooled OLS or Fixed effect or random effects model is better, we use LM test which says fixed effects model is appropriate. Further, to determine whether the Fixed effects model or random effect model is to be used. Hausman test is undertaken we conclude that random effect is appropriate (Table 8). The table test of model stationarity using Levin, Lin & Chu t test and model significance using Wald test is undertaken Table 9. both yield positive results. The table 11 & 12 explains check on autocorrelation of the data used probability values are greater than 0.05 thus we accept the null hypothesis that there is no autocorrelation. It means that the error terms are not equal to each other or correlated.

Table	5.	Descri	ptive	Sta	tistics
T CONTE	•••	Deseri	pure	Dia	cio cie o

Table 5. Descriptive Statistics							
	ER	HML	MOM	RM_RF	SMB		
Mean	-0.036	0.264	1.411	-0.111	-0.171		
Median	-0.161	-0.324	2.000	-0.114	-0.043		
Maximum	1.740	17.71	13.897	13.627	9.131		
Minimum	-1.77	-9.953	-20.970	-19.327	-13.695		
Std. Dev.	0.639	5.158	5.643	4.751	4.033		
Observations	1800	1800	1800	1800	1800		

Null Hypothesis:	thevariab	le has a i	mit root	,		
run mypoulesis.	At Level			1		
	<u>III Ee ver</u>	ER	HML	SMB	RMRF	MOM
With Constant	t-Statistic	0.0001	0.0000	0.0000	0.0000	0.0000
	Prob.	0.0001	0.0000	0.0000	0.0000	0.0000
		***	***	***	***	***
With Constant &						
Trend	t-Statistic	0.0000	0.0000	0.0000	0.0000	0.0000
	Prob.	0.0000	0.0000	0.0000	0.0000	0.0000
		***	***	***	***	***
Without Constant						
& Trend	t-Statistic	0.2667	0.0000	0.0000	0.0000	0.0000
	Prob.	0.2828	0.0000	0.0000	0.0000	0.0000
		***	***	***	***	***
	<u>At First</u> Diffe reno	ce				
			d(HML)	d(SMB)	d(RMRF)	d(MOM
With Constant	t-Statistic	0.0000	0.0000	0.0000	0.0000	0.0000
	Prob.	0.0000	0.0000	0.0000	0.0000	0.0000
		***	***	***	***	***
With Constant & Trend	t-Statistic	0.0000	0.0000	0.0000	0 0000	0.0000
Tiella	Prob.		0.0000	0.0000	0.0000	0.0000
	F 100.	***	***	***	***	***
Without Constant						
& Trend	t-Statistic	0.0000	0 0000	0 0000	0 0000	0.0000
	Proh.		0.0000	0.0000	0.0000	0.0000
	1100.	***	***	***	***	***
Notes:						

Table 7: Lagrange multiplie	er (LM) test for panel data

Null (no rand. effect)	Cross-section	Period	Both
Alternative	One-sided	One-sided	
Breusch-Pagan	70051.20	56.63865	70107.83
.0,	(0.0000)	(0.0000)	(0.0000)
Honda	264.6719	-7.525866	181.8297
10	(0.0000)	(1.0000)	(0.0000)
King-Wu	264.6719	-7.525866	247.9128
×	(0.0000)	(1.0000)	(0.0000)
SLM	273.4609	-7.341739	
	(0.0000)	(1.0000)	
GHM			70051.20

 Table 8: Test for determining Fixed Effect or Random

 Model

Correlated Random Effects - Hausman Test							
Equation: Untitled							
Test cross-section	andon	n effect	ts				
Test Summary		Chi-Sc	 Statistic 	Chi-Sq. d.f.	Prob.		
Cross-section rando	om	0.0000	000	4	1.0000		
* Cross-section tes	t varia	nce is i	invalid. Hau	sman statistic	set to		
zero.							
Cross-section rando	om eff	ects te	st compariso	ns:			
Variable	Fixed		Random	Var (Diff.)	Prob.		
HML	-0.00	0467	-0.000467	0.000000	0.3633		
MOM	MOM -0.000567 -0.000567 0.000000 0.3633						
SMB -0.000824 -0.000824 0.000000 0.363					0.3633		
RMRF	0.007	042	0.007042	0.000000	0.3633		



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Dependent Variable: ER								
Method: Panel Le								
Date: 10/25/22 Time: 11:42								
Sample: 1 1800								
Periods included:	120							
Cross-sections in	cluded: 15							
Total panel (balar	nced) observat	ions: 1800						
Variable	Coefficient	Std. Error	t-Statistic	Prob.				
HML	-0.000534	0.001717	-0.310637	0.7561				
МОМ	-0.002303	0.001521	-1.514002	0.1302				
RM_RF	-0.000585	0.001889	-0.309708	0.7568				
SMB	-0.001328	0.001766	-0.752032	0.4521				
С	-0.033115	0.007119	-4.651972	0.0000				
	Effects Specif	ication						
Cross-section fixe	ed (dummy va	riables)						
R-squared	0.810401	M ean deper	ndent var	-0.036542				
Adjusted R-								
squared	0.808484	S.D. depend	dent var	0.639428				
S.E. of								
regression	0.279830	Akaike info criterion 0.30122						
Sum squared								
resid	139.4607	Schwarz criterion 0.359237						
Log likelihood	-252.1060	Hannan-Qu		0.322642				
F-statistic	422.9160	Durbin-Wa	tson stat	1.851787				
Prob(F-statistic)	0.000000							

Hypothesis. We conclude that the Random effects model is appropriate.

 Table 9: Test of Normality

Group unit root test: Summary							
Series: RESIDUL, ROA, SALESG, SALESTO_ASSTS, SIZE,							
LEV, GROSS_MARGI	N, CASH, I	DEPENDEN	T_VARIA	BLE,			
REC_ASSTS							
Sample: 115							
Exogenous variables: In	dividual eff	ects					
Automatic selection of 1	naximum la	lgs					
Automatic lag length se							
Newey-West automatic	bandwidth	selection an	d Bartlett ke	ernel			
Balanced observations f	or each test		-				
			Cross-				
Method	Statistic	Prob.**	sections	Obs			
Null: Unit root (assume	s common u	nit root pro	cess)				
Levin, Lin & Chu t*	-12.6822	0.0000	9	126			
Null: Unit root (assume	s individual	unit root pr	ocess)				
Im, Pesaran and Shin							
W-stat	-9.31734	0.0000	9	126			
ADF - Fisher Chi-							
square	96.8730	0.0000	9	126			
PP - Fisher Chi-square	97.9244	0.0000	9	126			
** Probabilities for Fish	her tests are	computed u	sing an asyı	nptotic			
Chi							
-square distribution. All	other tests	assume asy i	mptotic nori	nality.			

Table 10: Test of Model Significance

Wald Test:				
Equation: Untitle	ed			
Test Statistic	Value	df	Probability	
F-statistic	9.827297	(8, 118)	0.0000	
Chi-square	78.61837	8	0.0000	
Null Hypothesis: C(1)=0,C(2)=0, C(3)=0, C(4)=0,C(5)=0,				

C(6)=0, C(7)=0, C(8) =0						
Null Hypothesis Su	Null Hypothesis Summary:					
Tabl	Table 11: Test of Autocorrelation					
Dependent Variable	: REC1					
Method: Panel Leas	t Squares					
Date: 06/04/22 Tim	e: 11:05					
Sample: 2012 2021						
Periods included: 10	0					
Cross-sections inclu	ded: 15					
Total panel (balance	ed) observati	ions: 150				
Variable	Coefficient	Std. Error	t-Statistic	Prob.		
SIZE	-0.027787	0.011265	-2.466586	0.0148		
ROA	-0.321599	0.054006	-5.954888	0.0000		
LEV	-0.274184	0.141312	-1.940277	0.0543		
CASH	-1.855129	34.47159	-0.053816	0.9572		
SALESG	-2.13E-05	0.000877 -0.024239		0.9807		
SALESASSTS	-0.068969	0.021523	-3.204472	0.0017		
GPROFIT	0.006460	0.099431	0.064969	0.9483		
RSESI	-0.088991	0.110969	-0.801942	0.4239		
С	0.627251	0.118270	5.303535	0.0000		
R-squared	0.392200	M ean depen	0.148416			
Adjusted R-squared	0.357715	S.D. depend	0.311277			
S.E. of regression	0.249465	Akaike info	0.119131			
Sum squared resid	8.774849	Schwarz cri	0.299769			
Log likelihood	0.065147	Hannan-Quinn criter. 0.1925				
F-statistic	11.37304	Durbin-Watson stat 1.08792				
Prob(F-statistic) 0.000000						

Table :12. Breusch-Godfrey Serial Correlation LM Test

Breusch-Godfrey Serial Correlation LM Test:					
F-statistic	1.088219	Prob. F (2,	0.4194		
Obs*R-squared	4.933287	Prob. Chi-	Square (2)	0.0849	
Test Equation:					
Dependent Variable:	RESID				
Method: Least Square	es				
Sample: 2 15					
Included observation:	s: 14				
Presample missing va	lue lagged re	esiduals set	to zero.		
Variable	Coefficient	Std. Error	t-Statistic	Prob.	
ROA	0.296944	0.443262	0.669907	0.5396	
SALESG	-5.579954	28.04898	-0.198936	0.8520	
SALESTO_ASSTS	0.021823	0.039859	0.547520	0.6131	
SIZE	0.003592	0.015319	0.234496	0.8261	
LEV	-0.183856	0.416595	-0.441330	0.6818	
GROSS_MARGIN	0.093205	0.456986	0.203956	0.8483	
CASH	15.78437	84.98009	0.185742	0.8617	
С	1.020003	5.548323	0.183840	0.8631	
RESID (-1)	-0.775631	0.537081	-1.444161	0.2222	
RESID (-2)	0.264941	0.731050	0.362412	0.7354	
R-squared	0.352378	Mean dependent var		5.20E-15	
Adjusted R-squared	-1.104773	S.D. dependent var		0.069272	
S.E. of regression	0.100498	Akaike info criterion		-1.581550	
Sum squared resid	0.040399	Schwarz criterion		-1.125081	
Log likelihood	21.07085	Hannan-Quinn criter.		-1.623805	
F-statistic	0.241827	Durbin-Watson stat		2.303277	
Prob(F-statistic)	0.964175				

Table 13 contains the summary statistics of both peer-firm and firm-specific characteristics with 150 firm-year



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observations. It provides details on all variables used in the analysis, firm-specific and averages for product-market peers alike. The cash holding ratio is less than 1%. The industry sales growth is around 23% with a gross profit of 70%. Table 14 reports the results from our analysis of the properties of the instrumental variable, where we regress peer idiosyncratic stock returns on the firm characteristics. In both specifications, using either contemporaneous or one-year lead values, none of the firm characteristics is correlated with the instrument except firm size while using Contemporaneous variables. Moreover, the F statistic of all firm characteristics is also insignificant, suggesting that peer idiosyncratic stock returns do not contain any significant information about the present or future observable determinants of firms' trade receivables. These results provide further assurance about the validity of the instrument used.

To identify whether trade credit is affected by peer firm, 2SLS regression approach, is adopted. Table 15 shows the results from both the first and second stages of regressions. Columns (1) and (2) present the estimates of our base regression of receivables to sales on the main independent variable, peer receivables to sales, while controlling for firm characteristics, peer averages, and year fixed effects. In the first-stage regression represented in (Column (1)), we find that the coefficient instrument, peer idiosyncratic stock returns, is significant, indicating that the average equity stock returns is strongly and positively related to average peer trade receivables. This result is consistent with prior studies that investment is positively correlated with stock returns. [11]. Fstatistic is quite large, suggesting that the instrument is likely to be valid. Second stage regression (Column (2)), the coefficient on peer receivables to sales (0.916) is positive and statistically significant at the 1% level, providing strong evidence of peer influence on trade credit policy[15]. In terms of the magnitude of peer effects, one standard deviation (0.046) increase in *peer receivables to sales* is associated with an increase of 0.43 in trade receivables.

	Table	13:	Descriptive	Analysis
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	Mean	Median
RECIVABLES_SALES	0.148	0.054
REC_ASSTS	0.064	0.052
PEER_RECSALES	0.862	0.139
PEER_REC_ASSTS	0.123	0.070
ROA	0.053	0.098
SIZE	9.765	9.689
LEV	0.372	0.385
CASH	0.000	0.000
SALESG	23.030	53.65
GROSS_MARGIN	0.706	0.748
SALESTO_ASSTS	1.308	0.998
RESIDUL	0.000	-0.010
PEER_SIZE	9.111	9.691
PEER_ROA	0.157	0.097
PEER_LEV	1.972	0.375
PEER_TANG	1.248	0.329
PEER_CASH	1.263	0.000

PEER_SALESG	18.316	35.20
PEER_GROSS_MARGIN	0.856	0.690
PEER_SALESTO_ASSTS	1.211	1.241
PEER_RESIDUL	2.618	-0.005

 Table 14. Estimating Peer Idiosyncratic stock returns

	Contemporaneous	One-year ahead
	independent variables	independent
SIZE	2.500581	2.588336
	0.0778	0.0643
TANG	-1.714858	-1.622727
	0.0888	0.1075
LEV	1.635852	0.520026
	0.1043	0.6041
CASH	-0.512784	-0.621295
	0.609	0.5357
ROA	-1.542226	-0.869282
	0.1255	0.3866
SALESG	3.609897	-0.457112
	0.0604	0.6485
SALESTO_ASSTS	0.881233	1.364708
	0.3799	0.1751
GROSS_MARGIN	0.443072	-0.255183
	0.6585	0.799
С	-2.443239	-2.422793
	0.0159	0.017
F-statistic	1.524793	0.400697
Durbin-Watson stat	2.408916	2.620207
Adjusted R-squared	0.071914	-0.109131

Table 15: Peer eff	ects in trade c	credit provision
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First stage		Second	
(1)	(2)	stage (1)	(2)
1.36	3.62		
		91.65	3.75
1.271	0.063	0.711	0.063
-14.052	0.002	12.62	0.001
7.105	0.030	6.545	0.012
-5.899	0.045	-4.563	0.023
-0.644	0.029	-1.204	0.029
18.890	0.002	12.26	0.000
-4.627	0.001	-5.187	0.000
0.0252	0.066	-0.5348	0.181
0.112	0.095	0.492	0.085
2.819	0.001	3.199	0.092
2.009	0.044	2.389	0.044
0.166	0.064	0.546	0.084
0.825	0.095	1.205	0.135
-0.034	0.096	0.346	0.097
-0.595	0.061	-0.215	0.041
43.004	0.020	22.023	0.000
0.67890			
315.8808			
0.000000			
	(1) 1.36 1.271 -14.052 7.105 -5.899 -0.644 18.890 -4.627 0.0252 0.112 2.819 2.009 0.166 0.825 -0.034 -0.595 43.004 0.67890 315.8808	1.36 3.62 1.271 0.063 -14.052 0.002 7.105 0.030 -5.899 0.045 -0.644 0.029 18.890 0.002 -4.627 0.001 0.0252 0.066 0.112 0.095 2.819 0.001 2.009 0.044 0.166 0.064 0.825 0.095 -0.034 0.096 -0.595 0.061 43.004 0.020 0.67890 315.8808 0.000000	(1) (2) stage (1) 1.36 3.62 91.65 1.271 0.063 0.711 -14.052 0.002 12.62 7.105 0.030 6.545 -5.899 0.045 -4.563 -0.644 0.029 -1.204 18.890 0.002 12.26 -4.627 0.001 -5.187 0.0252 0.066 -0.5348 0.112 0.095 0.492 2.819 0.001 3.199 2.009 0.044 2.389 0.166 0.064 0.546 0.825 0.095 1.205 -0.034 0.096 0.346 -0.595 0.061 -0.215 43.004 0.020 22.023 0.67890 315.8808 0.000000

Statistical significance at the 10 percent, 5 percent, and 1 percent levels, respectively.

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These findings collectively indicate that peer effects in trade credit provision are both statistically significant.

Our research adds to the practical knowledge of the growing literature on peer effects while formulating corporate financial policies. The study identifies that peer firm behaviour plays an important role in shaping the working capital policy. The implication of our research is that, managers should give predominant consideration to the peer firm behaviour while formulating trade credit decisions. The research can be further taken forward considering the level of competition, the size of firms and its impact on trade credit policies.

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